

## Section A: Q. 1 - Q. 10 Carry ONE mark each.

| Q.1 | Let $\left\{a_{n}\right\}_{n \geq 1}$ be a sequence of non-zero real numbers. Then which one of the <br> following statements is true? |
| ---: | :--- |
| (A) | If $\left\{\frac{a_{n+1}}{a_{n}}\right\}_{n \geq 1}$ sequence is a convergent sequence, then $\left\{a_{n}\right\}_{n \geq 1}$ is also a convergent |
| (B) | If $\left\{a_{n}\right\}_{n \geq 1}$ is a bounded sequence, then $\left\{a_{n}\right\}_{n \geq 1}$ is a convergent sequence |
| (C) | If $\left\|a_{n+2}-a_{n+1}\right\| \leq \frac{3}{4}\left\|a_{n+1}-a_{n}\right\|$ for all $n \geq 1$, then $\left\{a_{n}\right\}_{n \geq 1}$ is a Cauchy |
| sequence |  |
| (D) | If $\left\{\left\|a_{n}\right\|\right\}_{n \geq 1}$ is a Cauchy sequence, then $\left\{a_{n}\right\}_{n \geq 1}$ is also a Cauchy sequence |


| Q.2 | Let $f: \mathbb{R} \rightarrow \mathbb{R}$ be the function defined by |
| :--- | :--- |
|  | Then which one of the following statements is NOT true? |
| $\lim _{h \rightarrow 0} \frac{(x+h) \sin \left(\frac{1}{x}+h\right)-x \sin \frac{1}{x}}{h}$, |  |
| (A) | $f\left(\frac{2}{\pi}\right)=1$ |
| (B) | $f\left(\frac{1}{\pi}\right)=\frac{1}{\pi}$ |
| (C) | $f\left(-\frac{2}{\pi}\right)=-1$ |
| (D) | $f$ is not continuous at $x=0$ |



| Q. 4 | Let $A$ and $B$ be two events such that $0<P(A)<1$ and $0<P(B)<1$. <br> Then which one of the following statements is NOT true? |
| ---: | :--- |
| (A) | If $P(A \mid B)>P(A)$, then $P(B \mid A)>P(B)$ |
| (B) | If $P(A \cup B)=1$, then $A$ and $B$ cannot be independent |
| (C) | If $P(A \mid B)>P(A)$, then $P\left(A^{c} \mid B\right)<P\left(A^{c}\right)$ |
| (D) | If $P(A \mid B) \geq P(A)$, then $P\left(A^{c} \mid B^{c}\right)<P\left(A^{c}\right)$ |


| Q. 5 | If $M(t), t \in \mathbb{R}$, is the moment generating function of a random variable, then <br> which one of the following is NOT the moment generating function of any <br> random variable? |
| :--- | :--- |
| (A) | $\frac{5 e^{-5 t}}{1-4 t^{2}} M(t),\|t\|<\frac{1}{2}$ |
| (B) | $e^{-t} M(t), t \in \mathbb{R}$ |
| (C) | $\frac{1+e^{t}}{2\left(2-e^{t}\right)} M(t), t<\ln 2$ |
| (D) | $M(4 t), t \in \mathbb{R}$ |




| Q.8 | Let $X_{1}$ and $X_{2}$ be two independent and identically distributed discrete <br> random variables having the probability mass function <br> Then $P(x)= \begin{cases}\left(\frac{1}{2}\right)^{x}, & x=1,2,3, \ldots \\ 0, & \text { otherwise. }\end{cases}$ <br> (A) <br> $\frac{1}{256}$ <br> (B) <br> (C) $\frac{1}{512}$ <br> (D) <br> $\frac{9}{64}$ |
| :--- | :--- |


| Q.9 | Let $X_{1}, X_{2}, \ldots, X_{n}(n \geq 2)$ be a random sample from $\operatorname{Exp}\left(\frac{1}{\theta}\right)$ distribution, <br> where $\theta>0$ is unknown. If $\bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i}$, then which one of the following <br> statements is NOT true? |
| :--- | :--- |
| (A) | $\bar{X}$ is the uniformly minimum variance unbiased estimator of $\theta$ |$\quad$| (B) | $\bar{X}^{2}$ is the uniformly minimum variance unbiased estimator of $\theta^{2}$ |
| ---: | :--- |
| (C) | $\frac{n}{n+1} \bar{X}^{2}$ is the uniformly minimum variance unbiased estimator of $\theta^{2}$ |
| (D) | $\operatorname{Var}\left(E\left(X_{n} \mid \bar{X}\right)\right) \leq \operatorname{Var}\left(X_{n}\right)$ |


| Q.10 | Let $X_{1}, X_{2}, \ldots, X_{n}(n \geq 3)$ be a random sample from a $N\left(\mu, \sigma^{2}\right)$ <br> distribution, where $\mu \in \mathbb{R}$ and $\sigma>0$ are both unknown. Then which one of <br> the following is a simple null hypothesis? |
| ---: | :--- |
| (A) | $H_{0}: \mu<5, \sigma^{2}=3$ |
| (B) | $H_{0}: \mu=5, \sigma^{2}>3$ |
| (C) | $H_{0}: \mu=5, \sigma^{2}=3$ |
| (D) | $H_{0}: \mu=5$ |

Section A: Q. 11 - Q. 30 Carry TWO marks each.

| Q.11 | $\lim _{n \rightarrow \infty} \frac{6}{n+2}\left\{\left(2+\frac{1}{n}\right)^{2}+\left(2+\frac{2}{n}\right)^{2}+\cdots+\left(2+\frac{n-1}{n}\right)^{2}\right\}$ equals |
| :--- | :--- |
| (A) | 38 |
| (B) 36 |  |
| (C) 32 |  |
| (D) 30 |  |



| Q.13 | Let $f:[1,2] \rightarrow \mathbb{R}$ be the function defined by |
| :--- | :--- |
| $\qquad f(t)=\int_{1}^{t} \sqrt{x^{2} e^{x^{2}-1} d x .}$ |  |
| Then the arc length of the graph of $f$ over the interval $[1,2]$ equals |  |
| (A) | $e^{2}-\sqrt{e}$ |
| (B) | $e-\sqrt{e}$ |
| (C) | $e^{2}-e$ |
| (D) | $e^{2}-1$ |


| Q.14 | Let $F:[0,2] \rightarrow \mathbb{R}$ be the function defined by <br>  <br> where $[t]$ denotes the greatest integer less than or equal to $t . ~ T h e n ~ t h e ~ v a l u e ~$ <br> of the derivative of $F$ at $x=1$ equals <br> (A) <br> $e^{3}+2 e^{2}-e$ <br> (B) <br> $e^{3}-e^{2}+2 e$ <br> (C) <br> $e^{3}-2 e^{2}+e$ <br> (D) <br> $e^{3}+2 e^{2}+e$ |
| :--- | :--- |


| Q. 15 | Let the <br> have i <br> value | $\begin{aligned} -a y+z & =1 \\ -4 y+z & =-b \\ -y+2 z & =b+2 \end{aligned}$ <br> re $a$ and $b$ are rea |
| :---: | :---: | :---: |
| (A) | -11 |  |
| (B) | -10 |  |
| (C) | -13 |  |
| (D) | -14 |  |


| Q.16 | Let $A=\left(\begin{array}{lll}0 & 1 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 1\end{array}\right)$. Then the sum of all the elements of $A^{100}$ equals |
| :--- | :--- | :--- |
| (A) 101 |  |
| (B) 103 |  |
| (C) 102 |  |
| (D) 100 |  |


| Q.17 | Suppose that four persons enter a lift on the ground floor of a building. There <br> are seven floors above the ground floor and each person independently chooses <br> her exit floor as one of these seven floors. If each of them chooses the topmost <br> floor with probability $\frac{1}{3}$ and each of the remaining floors with an equal <br> probability, then the probability that no two of them exit at the same floor <br> equals |
| :--- | :--- |
| (A) | $\frac{200}{729}$ |
| (B) | $\frac{220}{729}$ |
| (C) | $\frac{240}{729}$ |
| (D) | $\frac{180}{729}$ |


| Q. 18 | A year is chosen at random from the set of years $\{2012,2013, \ldots, 2021\}$. <br> From the chosen year, a month is chosen at random and from the chosen month, <br> a day is chosen at random. Given that the chosen day is the $29^{\text {th }}$ of a month, <br> the conditional probability that the chosen month is February equals |
| :--- | :--- |
| (A) $\frac{279}{9965}$ |  |
| (B) $\frac{289}{9965}$ |  |
| (C) $\frac{269}{9965}$ |  |
| (D) $\frac{259}{9965}$ |  |


| Q.19 | Suppose that a fair coin is tossed repeatedly and independently. Let $X$ denote <br> the number of tosses required to obtain for the first time a tail that is <br> immediately preceded by a head. Then $E(X)$ and $P(X>4)$, respectively, <br> are |
| :--- | :--- |
| (A) | 4 and $\frac{5}{16}$ |
| (B) | 4 and $\frac{11}{16}$ |
| (C) | 6 and $\frac{5}{16}$ |
| (D) | 6 and $\frac{11}{16}$ |



| Q.21 | In a store, the daily demand for milk (in litres) is a random variable having <br> $\operatorname{Exp}(\lambda)$ distribution, where $\lambda>0$. At the beginning of the day, the store <br> purchases $c(>0)$ litres of milk at a fixed price $b(>0)$ per litre. The milk <br> is then sold to the customers at a fixed price $s(>b)$ per litre. At the end of <br> the day, the unsold milk is discarded. Then the value of $c$ that maximizes the <br> expected net profit for the store equals |
| :--- | :--- |
| (A) | $-\frac{1}{\lambda} \ln \left(\frac{b}{s}\right)$ |
| (B) | $-\frac{1}{\lambda} \ln \left(\frac{b}{s+b}\right)$ |
| (C) | $-\frac{1}{\lambda} \ln \left(\frac{s-b}{s}\right)$ |
| (D) | $-\frac{1}{\lambda} \ln \left(\frac{s}{s+b}\right)$ |


| Q. 22 | Let $X_{1}, X_{2}$ and $X_{3}$ be three independent and identically distributed random |
| ---: | :--- | :--- | :--- |
| (A) | $\frac{1}{6}$ |
| (B) | $\frac{1}{3}$ |
| (C) | $\frac{1}{8}$ |
| (D) | $\frac{1}{4}$ |


| Q.23 | Let $(X, Y)$ be a random vector having bivariate normal distribution with <br> parameters $E(X)=0, \operatorname{Var}(X)=1, E(Y)=-1, \operatorname{Var}(Y)=4$ and <br> $\rho(X, Y)=-\frac{1}{2}$, where $\rho(X, Y)$ denotes the correlation coefficient between <br> $X$ and $Y$. Then $P(X+Y>1 \mid 2 X-Y=1)$ equals |
| ---: | :--- |
| (A) | $\Phi\left(-\frac{1}{2}\right)$ |
| (B) | $\Phi\left(-\frac{1}{3}\right)$ |
| (C) | $\Phi\left(-\frac{1}{4}\right)$ |
| (D) | $\Phi\left(-\frac{4}{3}\right)$ |


| Q.24 | Let $\left\{X_{n}\right\}_{n \geq 1}$ be a sequence of independent and identically distributed random <br> variables having the common probability density function <br> If $\lim _{n \rightarrow \infty} P\left(\left\|\frac{1}{n} \sum_{i=1}^{n} X_{i}-\theta\right\|<\epsilon\right)=1$ for all $\epsilon>0$, then $\theta$ equals <br> $\frac{2}{x^{3}}$,$\quad x \geq 1$ |
| :--- | :--- |
| 0, | otherwise. |
| (A) | 4 |
| (B) 2 |  |
| (C) $\ln 4$ |  |
| (D) $\ln 2$ |  |


| Q.25 | Let $0.2,1.2,1.4,0.3,0.9,0.7$ be the observed values of a random sample of <br> size 6 from a continuous distribution with the probability density function |
| :--- | :--- |
| where $\theta>\frac{1}{2}$ is unknown. Then the maximum likelihood estimate and the |  |
| method of moments estimate of $\theta$, respectively, are |  |
| $\frac{1}{2 \theta-1}$, | $\frac{1}{2}<x \leq \theta$ |
| 0, | $0<x \leq \frac{1}{2}$ |
| (A) | $\frac{7}{5}$ and 2 |$\quad$| (B) $\frac{47}{60}$ and $\frac{32}{15}$ |
| :--- |
| (C) $\frac{7}{5}$ and $\frac{32}{15}$ |
| (D) $\frac{7}{5}$ and $\frac{47}{60}$ |


| Q.26 | For $n=1,2,3, \ldots$, let the joint moment generating function of $\left(X, Y_{n}\right)$ be |
| :--- | :--- |
|  | $M_{X, Y_{n}}\left(t_{1}, t_{2}\right)=e^{\frac{t_{1}^{2}}{2}}\left(1-2 t_{2}\right)^{-\frac{n}{2}}, t_{1} \in \mathbb{R}, t_{2}<\frac{1}{2}$ |
| If $T_{n}=\frac{\sqrt{n} X}{\sqrt{Y_{n}}}, n \geq 1$, then which one of the following statements is true? |  |$\quad$| (A) | The minimum value of $n$ for which $\operatorname{Var}\left(T_{n}\right)$ is finite is |
| ---: | :--- |
| (B) | $E\left(T_{10}^{3}\right)=10$ |
| (C) | $\operatorname{Var}\left(X+Y_{4}\right)=7$ |
| (D) | $\lim _{n \rightarrow \infty} P\left(\left\|T_{n}\right\|>3\right)=1-\frac{\sqrt{2}}{\sqrt{\pi}} \int_{0}^{3} e^{-\frac{t^{2}}{2}} d t$ |


| Q.27 | Let $X_{(1)}<X_{(2)}<\cdots<X_{(9)}$ be the order statistics corresponding to a <br> random sample of size 9 from $U(0,1)$ distribution. Then which one of the <br> following statements is NOT true? |
| ---: | :--- |
| (A) | $E\left(\frac{X_{(9)}}{1-X_{(9)}}\right)$ is finite |
| (B) | $E\left(X_{(5)}\right)=0.5$ |
| (C) | The median of $X_{(5)}$ is 0.5 |
| (D) | The mode of $X_{(5)}$ is 0.5 |


| Q. 28 | Let $X_{1}, X_{2}, \ldots, X_{16}$ be a random sample from a $N(4 \mu, 1)$ distribution and $Y_{1}, Y_{2}, \ldots, Y_{8}$ be a random sample from a $N(\mu, 1)$ distribution, where $\mu \in \mathbb{R}$ is unknown. Assume that the two random samples are independent. If you are looking for a confidence interval for $\mu$ based on the statistic $8 \bar{X}+\bar{Y}$, where $\bar{X}=\frac{1}{16} \sum_{i=1}^{16} X_{i}$ and $\bar{Y}=\frac{1}{8} \sum_{i=1}^{8} Y_{i}$, then which one of the following statements is true? |
| :---: | :---: |
| (A) | There exists a $90 \%$ confidence interval for $\mu$ of length less than 0 |
| (B) | There exists a $90 \%$ confidence interval for $\mu$ of length greater than 0.3 |
| (C) | $\left[\frac{8 \bar{X}+\bar{Y}}{33}-\frac{1.645}{2 \sqrt{66}}, \frac{8 \bar{X}+\bar{Y}}{33}+\frac{1.645}{2 \sqrt{66}}\right]$ is the unique $90 \%$ confidence interval for $\mu$ |
| (D) | $\mu$ always belongs to its 90\% confidence interval |


| Q. 29 | Let $X_{1}, X_{2}, X_{3}, X_{4}$ be a random sample from a distribution with the probability mass function $f(x)= \begin{cases}\theta^{x}(1-\theta)^{1-x}, & x=0,1 \\ 0, & \text { otherwise }\end{cases}$ <br> where $\theta \in(0,1)$ is unknown. Let $0<\alpha \leq 1$. To test the hypothesis $H_{0}: \theta=\frac{1}{2}$ against $H_{1}: \theta>\frac{1}{2}$, consider the size $\alpha$ test that rejects $H_{0}$ if and only if $\sum_{i=1}^{4} X_{i} \geq k_{\alpha}$, for some $k_{\alpha} \in\{0,1,2,3,4\}$. Then for which one of the following values of $\alpha$, the size $\alpha$ test does NOT exist? |
| :---: | :---: |
| (A) | $\frac{1}{16}$ |
| (B) | $\frac{1}{4}$ |
| (C) | $\frac{11}{16}$ |
| (D) | $\frac{5}{16}$ |


| Q.30 | Let $X_{1}, X_{2}, X_{3}, X_{4}$ be a random sample from a Poisson distribution with <br> unknown mean $\lambda>0$. For testing the hypothesis <br> $H_{0}: \lambda=1$ against $H_{1}: \lambda=1.5$, |
| :--- | :--- |
| let $\beta$ denote the power of the test that rejects $H_{0}$ if and only if |  |
| Then which one of the following statements is true? |  |$\quad$| (A) | $\beta>0.80$ |
| ---: | :--- |$\quad$| (B) $0.75<\beta \leq 0.80$ |
| :--- |
| (C) $0.70<\beta \leq 0.75$ |
| (D) $0.65<\beta \leq 0.70$ |

Section B: Q. 31 - Q. 40 Carry TWO marks each.

| Q.31 | Let $\left\{a_{n}\right\}_{n \geq 1}$ be a sequence of real numbers such that $a_{n}=\frac{1}{3^{n}}$ for all $n \geq 1$. <br> Then which of the following statements is/are true? <br> (A) <br> (B) <br> $\sum_{n=1}^{\infty}(-1)^{n+1} a_{n}$ is a convergent series <br> (C) <br> The radius of convergence of the power series $\sum_{n=1}^{\infty}\left(a_{1}+a_{2}+\ldots+a_{n}\right)$ is a convergent series $\frac{(-1}{3}$ |
| ---: | :--- |
| (D) | $\sum_{n=1}^{\infty} a_{n} \sin \frac{1}{a_{n}}$ is a convergent series |


| Q.32 | Let $f: \mathbb{R}^{2} \rightarrow \mathbb{R}$ be the function defined by <br> $f(x, y)=8\left(x^{2}-y^{2}\right)-x^{4}+y^{4}$. <br> Then which of the following statements is/are true? |
| :--- | :--- |
| (A) | $f$ has 9 critical points |
| (B) | $f$ has a saddle point at $(2,2)$ |$\quad$| (C) $f$ has a local maximum at $(-2,0)$ |
| ---: | :--- |


| Q.33 | If $n \geq 2$, then which of the following statements is/are true? |
| ---: | :--- |
| (A) | If $A$ and $B$ are $n \times n$ real orthogonal matrices such that <br> $\operatorname{det}(A)+\operatorname{det}(B)=0$, then $A+B$ is a singular matrix |
| (B) | If $A$ is an $n \times n$ real matrix such that $I_{n}+A$ is non-singular, then <br> $I_{n}+\left(I_{n}+A\right)^{-1}\left(I_{n}-A\right)$ |
| (C) | If $A$ is an $n \times n$ real skew-symmetric matrix, then $I_{n}-A^{2}$ is a non- |
| singular matrix |  |$\quad$| (D) | If $A$ is an $n \times n$ real orthogonal matrix, then $\operatorname{det}\left(A-\lambda I_{n}\right) \neq 0$ for all <br> $\lambda \in\{x \in \mathbb{R}: x \neq \pm 1\}$ |
| :--- | :--- |
|  |  |


| Q.34 | Let $\Omega=\{1,2,3, \ldots\}$ be the sample space of a random experiment and <br> suppose that all subsets of $\Omega$ are events. Further, let $P$ be a probability <br> function such that $P(\{i\})>0$ for all $i \in \Omega$. Then which of the following <br> statements is/are true? |
| ---: | :--- |
| (A) | For every $\epsilon>0$, there exists an event $A$ such that $0<P(A)<\epsilon$ |
| (B) | There exists a sequence of disjoint events $\left\{A_{k}\right\}_{k \geq 1}$ with $P\left(A_{k}\right) \geq 10^{-6}$ for |
| all $k \geq 1$ |  |$\quad$| (C) | There exists $j \in \Omega$ such that $P(\{j\}) \geq P(\{i\})$ for all $i \in \Omega$ |
| :--- | :--- |
| (D) | Let $\left\{A_{k}\right\}_{k \geq 1}$ be a sequence of events such that $\sum_{k=1}^{\infty} P\left(A_{k}\right)<\infty$. <br> Then for each $i \in \Omega$ there exists $N \geq 1$ <br> such that $i \notin \cup_{k=N}^{\infty} A_{k}$ |


| Q.35 | A university bears the yearly medical expenses of each of its employees up to a <br> maximum of Rs. 1000. If the yearly medical expenses of an employee exceed <br> Rs. 1000, then the employee gets the excess amount from an insurance policy <br> up to a maximum of Rs. 500. If the yearly medical expenses of a randomly <br> selected employee has $U(250,1750)$ distribution and $Y$ denotes the amount <br> the employee gets from the insurance policy, then which of the following <br> statements is/are true? |
| :--- | :--- |
| (A) | $E(Y)=\frac{500}{3}$ |
| (B) | $P(Y>300)=\frac{3}{10}$ |
| (C) | The median of $Y$ is zero |
| (D) | The quantile of order 0.6 for $Y$ equals 100 |


| Q.36 | Let $X$ and $Y$ be two independent random variables having $N\left(0, \sigma_{1}^{2}\right)$ and <br> $N\left(0, \sigma_{2}^{2}\right)$ distributions, respectively, where $0<\sigma_{1}<\sigma_{2}$. Then which of the <br> following statements is/are true? <br> (A) <br> $X+Y$ and $X-Y$ are independent <br> (B) <br> $2 X+Y$ and $X-Y$ are independent if $2 \sigma_{1}^{2}=\sigma_{2}^{2}$ |
| ---: | :--- |
| (C) | $X+Y$ and $X-Y$ are identically distributed |
| (D) | $X+Y$ and $2 X-Y$ are independent if $2 \sigma_{1}^{2}=\sigma_{2}^{2}$ |


| Q.37 | Let $(X, Y)$ be a discrete random vector. Then which of the following <br> statements is/are true? |
| ---: | :--- |
| (A) | If $X$ and $Y$ are independent, then $X^{2}$ and $\|Y\|$ are also independent. |
| (B) | If the correlation coefficient between $X$ and $Y$ is 1, then <br> $P(Y=a X+b)=1$ for some $a, b \in \mathbb{R}$ |
| (C) | If $X$ and $Y$ are independent and $E\left[(X Y)^{2}\right]=0$, then $P(X=0)=1$ or |
| $P(Y=0)=1$ |  |$\quad$ (D) | If $\operatorname{Var}(X)=0$, then $X$ and $Y$ are independent |
| :--- |


| Q.38 | Let $X_{1}, X_{2}$ and $X_{3}$ be three independent and identically distributed random <br> variables having $N(0,1)$ distribution. If <br>  <br> then which of the following statements is/are true? <br> $\left(X_{2}+X_{3}\right)^{2} \quad$ and $\quad V=\frac{2 X_{1}^{2}}{2 X_{1}^{2}+\left(X_{2}+X_{3}\right)^{2}}$ <br> (A) <br>  <br> (B) <br> has $F_{1,1}$ distribution and $V$ has $F_{1,2}$ distribution $F_{1,1}$ distribution and $V$ has $F_{2,1}$ distribution <br> (C) <br> (D) <br> $\frac{1}{2} V(1+U)$ and $V$ are independent $F_{2,3}$ distribution |
| :--- | :--- |



| Q.40 | Let $X_{1}, X_{2}, \ldots, X_{n}(n>1)$ be a random sample from a $N(\mu, 1)$ distribution, <br> where $\mu \in \mathbb{R}$ is unknown. Let $0<\alpha<1$. To test the hypothesis $H_{0}: \mu=0$ <br> against $H_{1}: \mu=\delta$, where $\delta>0$ is a constant, let $\beta$ denote the power of the <br> size $\alpha$ test that rejects $H_{0}$ if and only if $\frac{1}{n} \sum_{i=1}^{n} X_{i}>c_{\alpha}$, for some constant <br> $c_{\alpha}$. Then which of the following statements is/are true? |
| :--- | :--- |
| (A) | For a fixed value of $\delta, \beta$ increases as $\alpha$ increases |
| (B) | For a fixed value of $\alpha, \beta$ increases as $\delta$ increases |
| (C) | For a fixed value of $\delta, \beta$ decreases as $\alpha$ increases |
| (D) | For a fixed value of $\alpha, \beta$ decreases as $\delta$ increases |

## Section C: Q. 41 - Q. 50 Carry ONE mark each.

Let $\left\{a_{n}\right\}_{n \geq 1}$ be a sequence of real numbers such that $a_{1+5 m}=2$, $a_{2+5 m}=3, a_{3+5 m}=4, a_{4+5 m}=5, a_{5+5 m}=6, m=0,1,2, \ldots$. Then $\limsup _{\mathrm{n} \rightarrow \infty} a_{n}+\liminf _{\mathrm{n} \rightarrow \infty} a_{n}$ equals $\qquad$
Q. 42

Let $f: \mathbb{R} \rightarrow \mathbb{R}$ be a function such that

$$
20(x-y) \leq f(x)-f(y) \leq 20(x-y)+2(x-y)^{2}
$$

for all $x, y \in \mathbb{R}$ and $f(0)=2$. Then $f(101)$ equals $\qquad$


Let $X$ and $Y$ be two independent and identically distributed random variables having $U(0,1)$ distribution. Then $P\left(X^{2}<Y<X\right)$ equals $\qquad$ (round off to 2 decimal places)

| Q.45 | Consider a sequence of independent Bernoulli trials, where $\frac{3}{4}$ is the probability <br> of success in each trial. Let $X$ be a random variable defined as follows: If the <br> first trial is a success, then $X$ counts the number of failures before the next <br> success. If the first trial is a failure, then $X$ counts the number of successes <br> before the next failure. Then $2 E(X)$ equals |
| :--- | :--- |


| Q.46 | Let $X$ be a random variable denoting the amount of loss in a business. The <br> moment generating function of $X$ |
| :--- | :--- |
| $\qquad$$M(t)=\left(\frac{2}{2-t}\right)^{2}, t<2$. <br> If an insurance policy pays $60 \%$ of the loss, then the variance of the amount <br> paid by the insurance policy equals___ (round off to 2 decimal |  |

Q. 47

Let $(X, Y)$ be a random vector having the joint moment generating function

$$
M\left(t_{1}, t_{2}\right)=\left(\frac{1}{2} e^{-t_{1}}+\frac{1}{2} e^{t_{1}}\right)^{2}\left(\frac{1}{2}+\frac{1}{2} e^{t_{2}}\right)^{2}, \quad\left(t_{1}, t_{2}\right) \in \mathbb{R}^{2}
$$

Then $P(|X+Y|=2)$ equals $\qquad$ (round off to 2 decimal places)

Let $X_{1}$ and $X_{2}$ be two independent and identically distributed random variables having $\chi_{2}^{2}$ distribution and $W=X_{1}+X_{2}$. Then $P(W>E(W))$ equals $\qquad$ (round off to 2 decimal places)

| Q.49 | Let $2.5,-1.0,0.5,1.5$ be the observed values of a random sample of size 4 <br> from a continuous distribution with the probability density function |
| :--- | :--- |
| $\qquad f(x)=\frac{1}{8} e^{-\|x-2\|}+\frac{3}{4 \sqrt{2 \pi}} e^{-\frac{1}{2}(x-\theta)^{2}}, \quad x \in \mathbb{R}$, |  |
| where $\theta \in \mathbb{R}$ is unknown. Then the method of moments estimate of $\theta$ |  |
| equals | (round off to 2 decimal places) |

Let $X_{1}, X_{2}, \ldots, X_{25}$ be a random sample from a $N(\mu, 1)$ distribution, where $\mu \in \mathbb{R}$ is unknown. Consider testing of the hypothesis $H_{0}: \mu=5.2$ against $H_{1}: \mu=5.6$. The null hypothesis is rejected if and only if $\frac{1}{25} \sum_{i=1}^{25} X_{i}>k$, for some constant $k$. If the size of the test is 0.05 , then the probability of type-II error equals $\qquad$ (round off to 2 decimal places)

Section C: Q. 51 - Q. 60 Carry TWO marks each.
Q. 51 Let $f: \mathbb{R}^{2} \rightarrow \mathbb{R}$ be the function defined by $f(x, y)=x^{2}-12 y$.

If $M$ and $m$ be the maximum value and the minimum value, respectively, of the function $f$ on the circle $x^{2}+y^{2}=49$, then $|M|+|m|$ equals $\qquad$

| Q. 52 | The value of |
| :--- | :--- |

$$
\int_{0}^{2} \int_{0}^{2-x}(x+y)^{2} e^{\frac{2 y}{x+y}} d y d x
$$

equals $\qquad$ (round off to 2 decimal places)

# Q. 53 <br> Let $A=\left(\begin{array}{ccc}1 & -1 & 2 \\ -1 & 0 & 1 \\ 2 & 1 & 1\end{array}\right)$ and let $\left(\begin{array}{l}x_{1} \\ x_{2} \\ x_{3}\end{array}\right)$ be an eigenvector corresponding to 

 the smallest eigenvalue of $A$, satisfying $x_{1}^{2}+x_{2}^{2}+x_{3}^{2}=1$. Then the value of $\left|x_{1}\right|+\left|x_{2}\right|+\left|x_{3}\right|$ equals $\qquad$ (round off to 2 decimal places)| Q.54 | Five men go to a restaurant together and each of them orders a dish that is <br> different from the dishes ordered by the other members of the group. However, <br> the waiter serves the dishes randomly. Then the probability that exactly one of <br> them gets the dish he ordered equals <br> places) |
| :--- | :--- |

Q. 55 Let $X$ be a random variable having the probability density function

$$
f(x)= \begin{cases}a x^{2}+b, & 0 \leq x \leq 3 \\ 0, & \text { otherwise }\end{cases}
$$

where $a$ and $b$ are real constants, and $P(X \geq 2)=\frac{2}{3}$.
Then $E(X)$ equals $\qquad$ (round off to 2 decimal places)

| Q.56 | A vaccine, when it is administered to an individual, produces no side effects <br> with probability $\frac{4}{5}$, mild side effects with probability $\frac{2}{15}$ and severe side effects <br> with probability $\frac{1}{15}$. Assume that the development of side effects is independent <br> across individuals. The vaccine was administered to 1000 randomly selected <br> individuals. If $X_{1}$ denotes the number of individuals who developed mild side <br> effects and $X_{2}$ denotes the number of individuals who developed severe side <br> effects, then the coefficient of variation of $X_{1}+X_{2}$ equals <br> (round off to 2 decimal places) |
| :--- | :--- |

Q. 57

Let $\left\{X_{n}\right\}_{n \geq 1}$ be a sequence of independent and identically distributed random variables having $U(0,1)$ distribution. Let $Y_{n}=n \min \left\{X_{1}, X_{2}, \ldots, X_{n}\right\}$, $n \geq 1$. If $Y_{n}$ converges to $Y$ in distribution, then the median of $Y$ equals $\qquad$ (round off to 2 decimal places)

| Q. 58 | Let $X_{(1)}<X_{(2)}<X_{(3)}<X_{(4)}<X_{(5)}$ <br> random sample of size 5 <br> density function from a continuous distribution with the probability |
| :--- | :--- |
|  | $\quad f(x)= \begin{cases}\frac{1}{x^{2}}, & 1<x<\infty \\ 0, & \text { otherwise. }\end{cases}$ |
| Then the sum of all possible values of $r \in\{1,2,3,4,5\}$ for which $E\left(X_{(r)}\right)$ <br> is finite equals |  |



| Q. 60 | Let $X_{1}, X_{2}, \ldots, X_{9}$ be a random sample from a $N\left(\mu, \sigma^{2}\right)$ distribution, where |
| :--- | :--- |
|  | $\mu \in \mathbb{R}$ and $\sigma>0$ are unknown. Let the observed values of $\bar{X}=\frac{1}{9} \sum_{i=1}^{9} X_{i}$ |
| and $S^{2}=\frac{1}{8} \sum_{i=1}^{9}\left(X_{i}-\bar{X}\right)^{2}$ be 9.8 and 1.44, respectively. If the likelihood |  |
| ratio test is used to test the hypothesis $H_{0}: \mu=8.8$ against $H_{1}: \mu>8.8$, then |  |
| the $p$-value of the test equals _ (round off to 3 decimal places) |  |

